Uncertainty quantification for PDFs and machine learning

Pavel Nadolsky

Southern Methodist University

Based on

- arXiv:2306.03918, arXiv:2205.10444 [PRD 107 (2023) 3, 034008] with A. Cooper-Sarkar, <u>A. Courtoy</u>, T. Cridge, F. Giuli, L. Harland-Lang, <u>T.J. Hobbs</u>, <u>J. Huston</u>, X. Jing, R. S. Thorne, K. Xie, M. Yan, C.-P. Yuan
- studies with CTEQ-TEA, PDF4LHC working groups and Snowmass'21 EF06 Topical group





What determines the size of PDF uncertainties?



FIG. 2. Comparison of the PDFs at Q = 100 GeV. The PDFs shown are the N2LO sets of NNPDF4.0, CT18, MSHT20, ABMP16 with $\alpha_s(M_Z) = 0.118$, and ATLASpdf21. The ratio to the NNPDF4.0 central value and the relative 1σ uncertainty are shown for the gluon g, singlet Σ , total strangeness $s^+ = s + \bar{s}$, total charm $c^+ = c + \bar{c}$, up valence u^V and down valence d^V PDFs.

Figures from Snowmass'2021 whitepaper "Proton structure at the precision frontier", arXiv:2203.13923

Two types of modern error PDFs



Two powerful, complementary representations. Hessian PDFs can be converted into MC ones, and vice versa. P. Nadolsky, PhysTeV 2023 workshop

The tolerance puzzle

Why do groups fitting similar data sets obtain different PDF uncertainties?



The answer has direct implications for high-stake experiments such as *W* boson mass measurement, tests of nonperturbative QCD models and lattice QCD, high-mass BSM searches, etc.

Tolerances explained by epistemic uncertainties

Relative PDF uncertainties on the *gg* luminosity at 14 TeV in three PDF4LHC21 fits to the **identical** reduced global data set



While the fitted data sets are identical or similar in several such analyses, the differences in uncertainties can be explained by methodological choices adopted by the PDF fitting groups.

NNPDF3.1' and especially 4.0 (based on the NN's+ MC technique) tend to give smaller nominal uncertainties in data-constrained regions than CT18 or MSHT20

Epistemic uncertainties explain many such differences

Details in arXiv:2203.05506, arXiv:2205.10444

Aspects of PDF uncertainties

- 1. (Dis)agreement among fitted experimental data sets <
 - χ^2 tensions, L_2 sensitivity
- 2. Modeling of systematic uncertainties
- 3. Explicit and implicit priors in Hessian and NN/ML fits

associated with the **epistemic** uncertainty

explore using publicly available error PDFs and codes

Aspects of PDF uncertainties

- 1. (Dis)agreement among fitted experimental data sets
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associated with the epistemic uncertainty



Representative sampling



Epistemic PDF uncertainty...

...reflects **methodological choices** such as PDF functional forms or NN architecture and hyperparameters.

... can dominate the full uncertainty when experimental and theoretical uncertainties are small.

... is associated with the prior probability.

... can be estimated by **representative sampling** of the PDF solutions obtained with acceptable methodologies.

 \Rightarrow sampling over choices of experiments, PDF/NN functional space, models of correlated uncertainties...

 \Rightarrow in addition to sampling over data fluctuations



Components of PDF uncertainty



In each category, one must maximize

PDF fitting accuracy (accuracy of experimental, theoretical and other inputs)

F

PDF sampling accuracy

(adequacy of sampling in space of possible solutions)

Fitting/sampling classification is borrowed

from the statistics of large-scale surveys [Xiao-Li Meng, *The Annals of Applied Statistics*, Vol. 12 (2018), p. 685]

HEP is not alone

Various domains contend with **multi-dimensional non-probability samples**

Forecasting: presidential elections, financial markets, weather and climate, ... Meng, The Annals of Applied Statistics, 12(2), 685; Isakov and Kuriwaki, Harvard Data Science Review, 2(4), 2020

Political polling
M. R. Elliott, R. Valliant, Statistical Science, 32(2), 249 (2017)
M. A. Bailey, Polling at a Crossroads – Rethinking Modern Survey Research. Cambridge University Press, 2023

COVID-19 vaccination assessments and epidemiological studies *Bradley et al.*, <u>https://doi.org/10.1038/s41586-021-04198-4</u> *W. Dempsey, arXiv:2005.10425*

Clinical trials of medical treatments P. Msaouel, <u>https://doi.org/10.1080/07357907.2022.2084621</u>

Studies of biodiversity R. Boyd et al., <u>https://doi.org/10.1016/j.tree.2023.01.001</u>

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AI/ML techniques are superb for finding an excellent fit to data. Are these techniques adequate for uncertainty estimation [exploring all good fits]?

A common resampling procedure used by experimentalists and theorists:

- 1. Train a neural network model T_i with N_{par} (hyper)parameters on a randomly fluctuated replica of discrete data D_i . Repeat N_{rep} times. In a typical application: N_{par} > 10², N_{rep} < 10⁴.
- 2. Out of N_{rep} replicas T_i with "good" description of data [i.e., with a high likelihood $P(D_i|T_i) \propto e^{-\chi^2(D_i,T_i)/2}$], discard "badly behaving" (overfitted, not smooth, ...) replicas
- 3. Estimate the uncertainties of T_i using the remaining "well-behaved" replicas

Is this procedure rigorous? How many N_{rep} replicas does one need?

A likelihood-ratio test of NN models T_1 and T_2

From Bayes theorem, it follows that

 $\frac{P(T_2|D)}{P(T_1|D)} = \frac{P(D|T_2)}{P(D|T_1)} \times \frac{P(T_2)}{P(T_1)}$ $\equiv r_{\text{posterior}} \equiv r_{\text{likelihood}} \equiv r_{\text{prior}}$ $= a \text{leatory} \quad \text{epistemic + aleatory} \quad \text{probabilities}$ Suppose replicas T_1 and T_2 have the same r^2 [rul up to $r = \exp\left(\frac{\chi_1^2 - \chi_2^2}{2}\right) = 11$, but T_2

Suppose replicas T_1 and T_2 have the same $\chi^2 [r_{\text{likelihood}} = \exp\left(\frac{\chi_1^2 - \chi_2^2}{2}\right) = 1]$, but T_2 is disfavored compared to $T_1 [r_{\text{posterior}} \ll 1]$.

This only happens if $r_{\text{prior}} \ll 1$: T_2 is discarded based on its **prior** probability.

Epistemic PDF uncertainty is important in W boson mass and α_s measurements

ATLAS-CONF-2023-004

PDF-Set	p_{T}^{ℓ} [MeV]	$m_{\rm T}$ [MeV]	combined [MeV]
CT10	$80355.6^{+15.8}_{-15.7}$	$80378.1^{+24.4}_{-24.8}$	80355.8 ^{+15.7} -15.7
CT14	$80358.0^{+16.3}_{-16.3}$	80388.8 ^{+25.2} -25.5	$80358.4^{+16.3}_{-16.3}$
CT18	$80360.1^{+16.3}_{-16.3}$	$80382.2^{+25.3}_{-25.3}$	80360.4+16.3
MMHT2014	80360.3 ^{+15.9} -15.9	$80386.2^{+23.9}_{-24.4}$	80361.0 ^{+15.9} -15.9
MSHT20	80358.9 ^{+13.0} -16.3	$80379.4^{+24.6}_{-25.1}$	80356.3 ^{+14.6}
NNPDF3.1	$80344.7^{+15.6}_{-15.5}$	80354.3+23.6	80345.0 ^{+15.5} _15.5
NNPDF4.0	$80342.2^{+15.3}_{-15.3}$	80354.3 ^{+22.3} -22.4	$80342.9^{+15.3}_{-15.3}$

Table 2: Overview of fitted values of the *W* boson mass for different PDF sets. The reported uncertainties are the total uncertainties.

ATLAS-CONF-2023-015

The statistical analysis for the determination of $\alpha_s(m_Z)$ is performed with the xFitter framework [60]. The value of $\alpha_s(m_Z)$ is determined by minimising a χ^2 function which includes both the experimental uncertainties and the theoretical uncertainties arising from PDF variations:

$$\chi^{2}(\beta_{\exp},\beta_{th}) = \frac{\sum_{i=1}^{N_{data}} \left(\sigma_{i}^{\exp} + \sum_{j} \Gamma_{ij}^{\exp} \beta_{j,\exp} - \sigma_{i}^{th} - \sum_{k} \Gamma_{ik}^{th} \beta_{k,th}\right)^{2}}{\Delta_{i}^{2}} + \sum_{j} \beta_{j,\exp}^{2} + \sum_{k} \beta_{k,th}^{2}.$$

profiling of CT and MSHT PDFs requires to include a tolerance factor $T^2 > 10$ as in the ePump code

[T.J. Hou et al., <u>1912.10053</u>, Appendix F]

Also the next slide.

(1)

1. Tensions among experiments

Explore using the L_2 sensitivity

for Hessian PDFs

arXiv:2306.03918

by X. Jing, A. Cooper-Sarkar, A. Courtoy, T. Cridge, F. Giuli, L. Harland-Lang, T.J. Hobbs, J. Huston, P. N., R. S. Thorne, K. Xie, C.-P. Yuan



- Comparisons of strengths of constraints from individual data sets in 8 PDF analyses using the common L_2 sensitivity metric
- An interactive website (<u>https://metapdf.hepforge.org/L2/</u>) to plot such comparisons [2070 figures in total] 2023-06-15

Tolerance hypersphere in the PDF space

2-dim (i,j) rendition of N-dim (26) PDF parameter space

Hessian method: Pumplin et al., 2001

A symmetric PDF error for a physical

observable X is given by

$$\Delta X = \vec{\nabla} X \cdot \vec{z}_m = \left| \vec{\nabla} X \right|$$

$$= \frac{1}{2} \sqrt{\sum_{i=1}^{N} \left(X_i^{(+)} - X_i^{(-)} \right)^2}$$





Correlation cosine for observables X and Y: hep-ph/0110378 $\cos \varphi = \frac{\nabla X \cdot \nabla Y}{\Delta X \Delta Y} = \frac{1}{4\Delta X \Delta Y} \sum_{i=1}^{N} \left(X_i^{(+)} - X_i^{(-)} \right) \left(Y_i^{(+)} - Y_i^{(-)} \right) \xrightarrow{(b)}{Orthonormal eigenvector basis}$

LHC high-mass Drell-Yan process probes \bar{u} and \bar{d}



PDF correlations for A_{FB}

NEW: Courtoy, Fu, Hou, Hobbs, PN, Yuan, 2023

Drell-Yan forward-backward asymmetry can be sensitive to light sea and gluon for increasing $M_{\ell \bar{\ell}}$.



L_2 sensitivity, definition

 $S_{f,L_2}(E)$ for experiment *E* is the estimated $\Delta \chi_E^2$ for this experiment when a PDF $f_a(x_i, Q_i)$ increases by the +68% c.l. Hessian PDF uncertainty

Take $X = f_a(x_i, Q_i)$ or $\sigma(f)$; $Y = \chi_E^2$ for experiment E.



$$S_{f,L_2} \equiv \Delta Y(\vec{z}_{m,X}) = \vec{\nabla} Y \cdot \vec{z}_{m,X} = \vec{\nabla} Y \cdot \frac{\nabla X}{|\nabla X|} = \Delta Y \cos \varphi.$$

A fast version of the Lagrange Multiplier scan of χ_E^2 along the direction of $f_a(x_i, Q_i)$!

Lagrange multiplier scans on $g(x, M_H)$ at x = 0.01 and 0.3





What is the L₂ sensitivity?

- The L₂ sensitivity is a way of vizualizing the pulls of fitted experiments on the best-fit PDF $f_a(x, Q)$, for a particular parton flavor x, as a function of x and Q
 - or, when plotted for a PDF luminosity, as a function of the final-state mass M_X
- The best-fit value for a particular $f_a(x, Q)$ is determined by the sum of these pulls
- Both the L₂ and LM methods explore the parametric dependence of the χ^2 function in the vicinity of the global minimum
- The L₂ sensitivity streamlines comparisons among independent PDF analyses using published error PDFs
- The L₂ sensitivity has been used internally by CT (in CT18), by the PDF4LHC21 benchmarking group (to determine which data sets should be in the reduced PDF fit used for benchmarking), and now by ATLASpdf, CT, and MSHT

Estimated χ^2 pulls from experiments (L₂ sensitivity, T. J. Hobbs et al., arXiv:1904.00222)

CT18 NNLO, g(x, 100 GeV)



CT18 NNLO, gluon at Q=100 GeV

15 core-minutes



Experiments with large $\Delta \chi^2 > 0$ [$\Delta \chi^2 < 0$] pull g(x, Q) in the negative [positive] direction at the shown *x*

Estimated χ^2 pulls from experiments (L₂ sensitivity, T. J. Hobbs et al., arXiv:1904.00222)

CT18 NNLO, g(x, 100 GeV)



CT18 NNLO, gluon at Q=100 GeV

Most sensitive experiments						
253 ATL8ZpTbT	109 cdhswf3					
542 CMS7jtR7y6T	110 ccfrf2.mi					
544 ATL7jtR6uT	147 Hn1X0c					
545 CMS8jtR7T	<mark>204</mark> e866ppxf					
160 HERAIpII	504 cdf2jtCor2					
101 BcdF2pCor						
102 BcdF2dCor						
<u>+08</u> cdhswf2						

Note opposite pulls (tensions) in some x ranges between HERA I+II DIS (ID=160); CDF (504), ATLAS 7 (544), CMS 7 (542), CMS 8 jet (545) production; E866pp DY (204); ATLAS 8 Z pT (253) production; BCDMS and CDHSW DIS

CT18 NNLO g(x, 100 GeV)



2023-06-15





Can also look at L₂ for 2 GeV



Examine the impact of each experiment on the different PDFs





since first derivative of χ^2 vanishes at the global minimum, the sum of the L₂ sensitivities must be zero within uncertainties

$$\sum_{E} S_{f,L_2} \ll T^2 < \sum_{E} |S_{f,L_2}|$$



2023-06-15

MSHT20 and CT18

Note importance of ATLAS Z p_T data (also, Z p_T data poorly fit at NNLO)

2023-06-15



MSHT20 NNLO and aN3LO



MSHT20 NNLO and aN3LO



at aN3LO, the two experiments now on same side; aN3LO needed for HERA

ATLASpdf21

ATLAS PDF fits are based on a more limited set of data, with HERA inclusive as the backbone

series of sequential PDF fits adding W/Z data,ttbar, W/Z+jets, inclusive jets and photon ratio data

full information on correlated systematic sources of uncertainty used (not available to to other PDF fits)

ID	Data set	\sqrt{s} [TeV]	Luminosity $[fb^{-1}]$	Decay channel	Observables entering the fit
160	HERA inclusive DIS [26]	Varied	Varied		Reduced cross sections
68	Inclusive $W, Z/\gamma^*$ [27]	7	4.6	e, μ combined	$\eta_\ell \; (W), y_Z \; (Z)$
89	Inclusive Z/γ^* [28]	8	20.2	e, μ combined	$\cos \theta^*$ in bins of $y_{\ell\ell}, m_{\ell\ell}$
86	Inclusive W [29]	8	20.2	μ	η_{μ}
56	$W^{\pm} + \text{jets}$ [30]	8	20.2	e	p_{T}^W
50	Z + jets [31]	8	20.2	e	$p_{\rm T}^{\rm jet}$ in bins of $ y^{\rm jet} $
7	$t\bar{t}$ [32, 33]	8	20.2	lepton + jets, dilepton	$m_{tar{t}},p_{\mathrm{T}}^t,y_{tar{t}}$
8	$t\bar{t}$ [34]	13	36	lepton + jets	$m_{tar{t}},p_{\mathrm{T}}^{t},y_{t},y_{tar{t}}^{\mathrm{b}}$
9	Inclusive isolated γ [35]	8, 13	$20.2, \ 3.2$	-	E_{T}^{γ} in bins of η^{γ}
10	Inclusive jets [36–38]	7, 8, 13	4.5, 20.2, 3.2	-	$p_T^{ m jet}$ in bins of $ y^{ m jet} $



Impact of addition of W/Z, ttbar data to HERA inclusive





FIG. 22. Sensitivities for ATLAS 8 TeV (left column) and CMS 8 TeV (right column) $t\bar{t}$ data sets in the CT18 (upper row) and MSHT20 (middle and lower rows) analyses at Q = 100 GeV.

Many more such comparisons on https://metapdf.hepforge.org/L2/

2. Implementations of systematic uncertainties

Explore using a hopscotch scan for MC PDFs

arXiv:2205.10444 [PRD 107 (2023) 3, 034008]

by A. Courtoy, J. Huston, P. N., K. Xie, M. Yan, C.-P. Yuan

Goodness-of-fit functions in PDF analyses

Analysis	χ ² prescription to fit PDFs	χ^2 prescription to compare PDFs	Comments
HERAPDF	HERA	HERA	
СТ	Extended T +addl. prior	Extended <i>T</i> , Experimental	
MSHT'20	Т	Т	
NNPDF4.0	<i>t</i> ₀ +addl. prior with fluctuated cross-sampled data	Experimental or t ₀ with unfluctuated full data	<i>t</i> ₀ prescription has pre- and post-NNPDF3.0 versions
Hopscotch'2022	N/A	Experimental or t_0 [2022] with unfluctuated data	

Different prescriptions reflect modeling of additive and multiplicative systematic errors in covariance matrices. Neither prescription is complete because of the bias-variance dilemma. The χ^2 definition affects the PDF uncertainty.

Two forms of χ^2 in PDF fits

1. In terms of nuisance parameters $\lambda_{\alpha,exp}$

$$\chi^2 = \sum_{i=1}^{N_{pt}} \frac{\left[D_i + \sum_{\alpha} \beta_{i,\alpha}^{\exp} \lambda_{\alpha,\exp} - T_i\right]^2}{s_i^2} + \sum_{\alpha} \lambda_{\alpha,\exp}^2$$

2. In terms of the covariance matrix

$$\chi^{2} = \sum_{i,j}^{N_{pt}} (T_{i} - D_{i})(\operatorname{cov}^{-1})_{ij} (T_{j} - D_{j})$$
$$(\operatorname{cov})_{ij} \equiv s_{i}^{2} \delta_{ij} + \sum_{\alpha=1}^{N_{\lambda}} \beta_{i,\alpha} \beta_{j,\alpha}, \qquad \qquad \beta_{i,\alpha} = \sigma_{i,\alpha} X_{i},$$

 D_i , T_i , s_i are the central data, theory, uncorrelated error $\beta_{i,\alpha}$ is the correlation matrix for N_{λ} nuisance parameters.

Experiments publish $\sigma_{i,\alpha}$. To reconstruct $\beta_{i,\alpha}$, we need to decide on the normalizations X_i . Possible choices:

a.
$$X_i = D_i$$
 : "**exp**erimental scheme"; can result in a bias
b. X_i = fixed or varied T_i : " t_0 , T, extended T schemes"; can result in (different) biases

Systematic uncertainties and the bias-variance dilemma

$$\chi^{2} = \sum_{i,j}^{N_{pt}} (T_{i} - D_{i})(\text{cov}^{-1})_{ij} (T_{j} - D_{j}) \qquad (\text{cov})_{ij} = s_{i}^{2} \delta_{ij} + \sum_{\alpha=1}^{N_{\lambda}} \beta_{i,\alpha} \beta_{j,\alpha}$$

$$\beta_{i,\alpha} = \sigma_{i,\alpha} X_i$$

 D_i , T_i , s_i are the central data, theory, uncorrelated error

 $\beta_{i,\alpha} \equiv \sigma_{i,\alpha} \hat{X}_i$ is the correlation matrix for N_{λ} nuisance parameters. Experiments publish $\sigma_{i,\alpha}$.

The "truth" normalizations \hat{X}_i in the experiment are of order T_i or D_i . { \hat{X}_i } are learned as a model { X_i } together with PDFs f and theory { $T_i(f)$ }. For example, we can sample as $X_i = a_i D_i + b_i T_i$, with free $0 \le a_i, b_i \le 1$.

Mean variation δ_X^2 of the model from truth on an ensemble of replicas, for data point *i*:

$$\delta_X^2 \equiv \left\langle \left(X_i - \hat{X}_i\right)^2 \right\rangle = \underbrace{\left\langle \left(\hat{X}_i - \langle X_i \rangle\right)^2 \right\rangle}_{\text{model bias}} + \underbrace{\left\langle (X_i - \langle X_i \rangle)^2 \right\rangle}_{\text{variance}} = \underbrace{\left\langle \left(\hat{X}_i - \langle X_i \rangle\right)^2 \right\rangle}_{\text{model bias}} - \underbrace{\left\langle (D_i - \langle X_i \rangle)^2 \right\rangle}_{\text{data bias}} + \underbrace{\left\langle (D_i - X_i)^2 \right\rangle}_{\chi^2(D_i, T_i)}$$

Experimental definition, $X_i = D_i$: $\langle (X_i - \hat{X}_i)^2 \rangle = (\hat{X}_i - D_i)^2 \equiv \delta_D^2$

$$t_0$$
 definition, $X_i = t_{0i}$: $\left\langle \left(X_i - \hat{X}_i\right)^2 \right\rangle = \left(\hat{X}_i - t_{0i}\right)^2 \equiv \delta_{t_0}^2$

In general, not enough information to compare δ_D and δ_{t_0}

Smoothing of *K*-factors

An analogous **bias-variance tradeoff** arises during smoothing of MC integration errors for *K*-factor tables

A smoother curve for theory reduces the χ^2 for the jet data, but the best-fit result retains some dependence on the fitted functional form

This dependence can be conservatively estimated by including an uncorrelated MC integration error (both the statistical error provided with the NNLO calculation and the epistemic error due to the functional form)





A hopscotch study of NN4.0 PDFs: look for signs of epistemic uncertainty

- **1.** Scan the quasi-Gaussian χ^2 dependence along 50 Hessian EV directions
- **2. Sample** with high density a few EV directions that drive the specific PDF uncertainty



Hopscotch scan+sampling of PDF parametrizations



Nominal NN4.0 Hessian or MC 68%cl

least as large as shown



The hopscotch scans: NNPDF4.0 vs CT18 uncertainties



The ellipses are projections of 68% c.l. ellipsoids in N_{par} -dim. spaces

 $N_{par} = 28$ and 50 for CT18 and NNPDF4.0 Hessian PDFs

Hopscotch scans realize the likelihood-ratio test



According to the LR test, the NN4.0 analysis discards PDFs in the green and blue regions based on the prior probabilities and differences in the likelihood definitions – both associated with prior terms

The allowed regions will change for the other acceptable χ^2 definitions, which exist in reflection of the biasvariance dilemma

PDF wish list for systematic uncertainties A proposal

- 1. More complete representations for experimental likelihoods that do not need reverse engineering
- 2. Agreed-upon nomenclature for leading syst. sources
- Is reducing dimensionality of published correlation matrices advisable? Is their a standard for it? E.g., fewer nuisance parameters; collect less relevant/certain nuisance parameters into one uncorrelated error; etc.
- 4. Mathematical consistency of covariance/correlation matrices (see Z. Kassabov et al.)
- 5. How do different implementations of syst. errors affect pulls on PDFs? L_2 sensitivities to nuisance parameters

6. ...

Final remarks

Epistemic uncertainty (due to parametrization, methodology, parametrization/NN architecture, smoothness, data tensions, model for syst. errors, ...) is increasingly important in NNLO global fits as experimental and theoretical uncertainties decrease

Nominal PDF uncertainties in high-stake measurements (ATLAS W mass, Higgs cross sections...) thus should be tested for *control of tensions* and *robustness of sampling over acceptable methodologies.*

Smoothness of Hessian and NN PDFs is another such aspect associated with the prior that should be explored.

Such tests can be done outside of the PDF fits.

Tools for such studies exist using published PDFs and codes: L_2 sensitivities and hopscotch scans.

This is also necessary for combination of PDFs including data correlations [LHC EW, Jet & Vector boson WGs, <u>https://tinyurl.com/4wcnd8xn</u>; <u>https://tinyurl.com/2p8d8ba3</u>; <u>https://tinyurl.com/2p8tcn5b</u>; Ball, Forte, Stegeman, arXiv:2110.08274].

The ambiguity in NNLO PDFs due to the χ^2 definition is significant. Must consider better formats to propagate experimental likelihoods into the PDF uncertainties. [See also Cranmer, Prosper, et al., arXiv:2109.04981].

Uncertainty quantification, a challenge for AI, As we try to analyze PDFs and understand "why". With machine learning methods we strive To make sense of the data and derive.

But uncertainty presents a hurdle As we try to make predictions and be certain. It's a challenge that we must face As we work to improve our models with grace.

Parton distributions, oh how they vex As we try to understand their complex effects. But still we persist, for we must know The secrets that uncertainty has yet to show.

Early Microsoft Bing

Backup

Augmented likelihood for PDFs with global tolerance

1. Start by defining the correspondence between $\Delta \chi^2$ and cumulative probability level: 68% c.l. $\Leftrightarrow \Delta \chi^2 = T^2$. 2. Write the **augmented** likelihood density for this definition:

 $P(D_i|T_i) \propto e^{-\chi^2/(2T^2)}$

3. When profiling 1 new experiment with the prior imposed on PDF nuisance parameters $\lambda_{\alpha,th}$:

$$\chi^{2}(\vec{\lambda}_{exp},\vec{\lambda}_{th}) = \sum_{i=1}^{N_{pt}} \frac{\left[D_{i} + \sum_{\alpha} \beta_{i,\alpha}^{exp} \lambda_{\alpha,exp} - T_{i} - \sum_{\alpha} \beta_{i,\alpha}^{th} \lambda_{\alpha,th}\right]^{2}}{s_{i}^{2}} + \sum_{\alpha} \lambda_{\alpha,exp}^{2} + \sum_{\alpha} T^{2} \lambda_{\alpha,th}^{2}. \qquad \beta_{i,\alpha}^{th} = \frac{T_{i}(f_{\alpha}^{+}) - T_{i}(f_{\alpha}^{-})}{2},$$

$$new \text{ experiment} \qquad priors \text{ on expt. systematics} and PDF \text{ params}$$
4. Alternatively, we can reparametrize $\chi^{2'} \equiv \chi^{2}/T^{2}$, so that 68% c.l. $\Leftrightarrow \Delta \chi^{2'} = 1$. We have
$$P(D_{i}|T_{i}) \propto e^{-\chi^{2'/2}}$$

$$\chi^{2'}(\vec{\lambda}_{exp}, \vec{\lambda}_{th}) = \sum_{i=1}^{N_{pt}} \frac{\left[D_{i} + \sum_{\alpha} \beta_{i,\alpha}^{exp} \lambda_{\alpha,exp} - T_{i} - \sum_{\alpha} \beta_{i,\alpha}^{th} \lambda_{\alpha,th}\right]^{2}}{s_{i}^{2} T^{2}} + \sum_{\alpha} \lambda_{\alpha,exp}^{2} + \sum_{\alpha} \lambda_{\alpha,exp}^{2}.$$

5. Inconsistent redefinitions:

$$\chi^{2}(\vec{\lambda}_{exp},\vec{\lambda}_{th}) = \sum_{i=1}^{N_{pt}} \frac{\left[D_i + \sum_{\alpha} \beta_{i,\alpha}^{exp} \lambda_{\alpha,exp} - T_i - \sum_{\alpha} \beta_{i,\alpha}^{th} \lambda_{\alpha,th}\right]^2}{s_i^2} + \sum_{\alpha} \lambda_{\alpha,exp}^2 + \sum_{\alpha} \lambda_{\alpha,th}^2. \qquad \text{and } P(D_i|T_i) \propto e^{-\chi^2/2} + \sum_{\alpha} \lambda_{\alpha,th}^2 + \sum_{\alpha} \lambda_{$$

Why augmented likelihood?

The term is accepted in lattice QCD to indicate that the log-likelihood contains prior terms

$$\chi^{2}(\vec{\lambda}_{\exp},\vec{\lambda}_{th}) = \sum_{i=1}^{N_{pt}} \frac{\left[D_{i} + \sum_{\alpha} \beta_{i,\alpha}^{\exp} \lambda_{\alpha,\exp} - T_{i} - \sum_{\alpha} \beta_{i,\alpha}^{th} \lambda_{\alpha,th}\right]^{2}}{s_{i}^{2}} + \sum_{\alpha} \lambda_{\alpha,\exp}^{2} + \sum_{\alpha} T^{2} \lambda_{\alpha,th}^{2}.$$
new experiment
priors on expt. systematics
and PDF params

After minimization w.r.t. to $\lambda_{\alpha,exp}$, $\lambda_{\alpha,th}$, the prior terms are **hidden** inside the covariance matrix:

$$\chi^{2} = \sum_{i,j}^{N_{pt}} (T_{i} - D_{i}) (\text{cov}^{-1})_{ij} (T_{j} - D_{j})$$

The usual χ^2 definition therefore contains a **prior** component, which may be handled differently by the various groups

Computing uncertainty ΔX in the NNPDF analysis

1. By unweighted averaging of predictions for 100 (or 1000) MC replicas:

$$\langle X \rangle = \frac{1}{N_{rep}} \sum_{i=1}^{N_{rep}} X_i; \quad \Delta X^2 = \langle (X - \langle X \rangle)^2 \rangle$$

NNPDF calls it "**importance sampling**". The MC replicas are distributed according to the fluctuated data [Ball:2011gg] using the same training algorithm.

This estimates the **aleatory** uncertainty for a given methodology.

Replica 0 is the mean of 1000 MC replicas; has better unfluctuated χ^2 than MC replicas.

2. Using $N_{eig} = 50$ Hessian PDFs.

$$\Delta X^2 = \sum_{i=1}^{N_{eig}} (X_i - X_0)^2 \,.$$

NNPDF4.0 MC and Hessian uncertainties are in a good agreement.



The hopscotch scan counterbalances the bias of the nominal replica ensemble

6.2 Creating a less biased sub-sample

The basic idea is to use such partial information about the selection bias to design a *biased* subsampling scheme to *counterbalance* the bias in the original sample, such that the resulting sub-samples have a *high likelihood* to be less biased than the original sample from our target population. That is, we create a sub-sampling indicator S_I , such that with high likelihood, the correlation between S_IR_I and G_I is reduced, compared to the original $\rho_{R,G}$, to such a degree that it will compensate for the loss of sample size and hence reduce the MSE of our estimator (e.g., the sample average). We say with *high likelihood*, in its non-technical meaning, because without full information on the response/recording mechanism, we can never guarantee such a counterbalance sub-sampling (CBS) would always do better. However, with judicious execution, we can reduce the likelihood of making serious mistakes.

X.-L. Meng, Survey Methodology, Catalogue 12-001-X, vol. 48 (2022), #2



Why doesn't NNPDF4.0 find HS solutions?



NNPDF authors find that some HS replicas fail the initial-stage overfitting test (M. Ubiali, HP2 2022 workshop, Durham, 2022-09-22)



xg (x,Q) at Q=1.7 GeV (sym. err) NNPDF4.0 NNLO 68% (solid), alt. $(\Delta \chi^2)_{t0}$ =0 (dashed)



HS solutions have much lower χ^2 than NN MC replicas. HS PDFs are outside the 50-dim neighborhood of NN replica 0. We do not see evidence of "overfitting" according to CT18 criteria.





Search docs

Getting started

Fitting code: n3fit

Code for data: validphys

/ Chi square figures of merit

Chi square figures of merit

Within the NNPDF methodology various figures of merit are used, each of which can be used in different situations. To avoid confusion, it is important to understand the differences between the various figures of merit, and to understand which definition we are referring to in a given context. In particular, it is worth stressing that whenever a figure of merit is discussed, the t_0 method (discussed below) applies.

Note

From NNPDF2.0 onwards the t_0 formalism has been used to define the figure of merit used during the fitting of the PDFs.

Note

The t_0 method is **not** used by default in other validphys applications, and instead the default is to compute the experimental χ^2 . To compute $\chi^2_{t_0}$, users need to specify

use_t0: True
t0pdfset: <Some LHAPDF set>

in the relevant namespace. This will instruct actions such as validphys.results.dataset chi2 table() to compute the t_0 estimator.

https://docs.nnpdf.science/figuresofmerit/index.html, accessed on 2023-03-28

Hopscotch NN4.0 replicas

LHAPDF6 grids available at https://ct.hepforge.org/PDFs/2022hopscotch/



Scans of the log-likelihood in EV directions 25 and 33



PRIOR PROBABILITY IN PDF FITS

✓ PDF fitting example of inverse problem: aim to find a posterior probability of **f** given the data **D**.

✓ Parametrization of PDFs: finite-dimensional problem.

 $f(x) \approx \tilde{f}(x,\theta) \in \mathcal{F}$

✓ The posterior probability for the parametrization depends on both the figure of merit that maximises the data likelihood given the parameters and on prior probability *H*.

(M. Ubiali, HP2 2022 workshop, Durham, 2022-09-22)